

Goodness-of-Fit Test: Testing a Population's Distribution (Section 12.1)

Here, our null hypothesis will be that the population follows a certain distribution. The alternative hypothesis is that it does *not*.

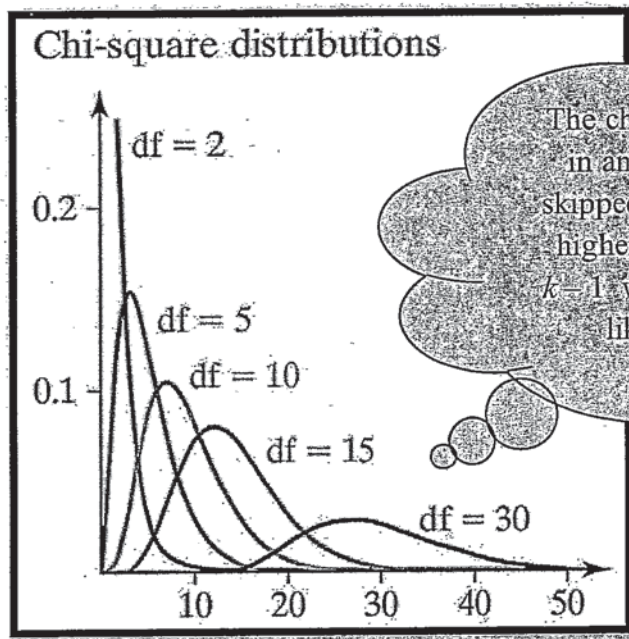
The M&Ms company wants to ensure that the distribution of colors in every bag roughly follows this distribution. They sample a random bag (counting each color) and use that data to perform a goodness-of-fit test to see if the distribution is what they want. Let's see how that works.

| Color | Percent |
|--------|---------|
| brown | 13% |
| yellow | 14% |
| red | 13% |
| orange | 20% |
| blue | 24% |
| green | 16% |

Definition: Goodness-of-Fit Test: A goodness-of-fit test is an inferential procedure used to determine whether a frequency distribution follows a specific distribution. We use the **Chi-Square (χ^2) Distribution**.

Characteristics of the Chi-Square Distribution:

1. It is *not* symmetric.
2. Its shape depends on the degrees of freedom, just like Student's t -distribution.
3. As the number of degrees of freedom increases, it becomes more symmetric, as illustrated in the figure below.
4. The values of χ^2 are non-negative (greater than or equal to 0).



The chi-square distribution was introduced in an earlier section of the book that we skipped. Notice how its shape changes with higher degrees of freedom (which will be $k-1$ where k is the number of categories, like colors in the above example).

Again, we find critical values or use P -values. Table VIII will give us those critical values.

In practice, what we will do is compare the actual observations (from the sample) to what we would *expect* them to be, given that the null hypothesis is true. If a significant difference exists between the observed and expected counts, we have evidence against the null hypothesis. So, how do we do that?

Expected Counts:

This is related to the expected values for binomial distributions and so will look familiar.

Suppose that there are n independent trials of an experiment (or individuals in a survey) with $k \geq 3$ mutually exclusive possible outcomes (categories). Let p_1 represent the (assumed) probability of observing the first outcome, p_2 represent the (assumed) probability of observing the second outcome, and so on. Then the **expected counts** for each possible outcome are given by $E_i = \mu_i = np_i$ for $i = 1, 2, \dots, k$.

These p_i probabilities are those *assumed* in the null hypothesis.

Goodness-of-Fit Test Statistic:

Let O_i represent the *observed* count of category i (data from sample or experiment) and E_i represents the *expected* count of category i . The number of categories is denoted by k and the number of independent trials of the experiment is n . Then

$$\chi^2_0 = \sum_{i=1}^k \frac{(O_i - E_i)^2}{E_i}$$

We use sigma notation here to indicate addition of these terms for $i = 1, 2, \dots, k$.

approximately follows the chi-square distribution with $k - 1$ degrees of freedom, provided that

1. all expected counts are greater than or equal to 1 (all $E_i \geq 1$), and
2. no more than 20% of the expected counts are less than 5.

Again, we calculate $E_i = \mu_i = np_i$ for $i = 1, 2, \dots, k$.

If these are *not* met, one option is to combine low-frequency categories into a single category.

The Goodness-of-Fit Test:

Step 1: Determine the null and alternative hypotheses.

H_0 : The random variable follows a certain distribution.

H_1 : The random variable does *not* follow this distribution.

This is the only
option for step 1.

Step 2: Select a level of significance, α , depending on the seriousness of making a Type I error.

Step 3a: Calculate the expected counts, $E_i = \mu_i = np_i$ for $i = 1, 2, \dots, k$. Recall k is the number of categories, n is the number of trials, and p_i is the *assumed* probability of the i^{th} category from the null hypothesis.

Step 3b: Verify that

1. all expected counts are greater than or equal to 1 (all $E_i \geq 1$), and
2. no more than 20% of the expected counts are less than 5.

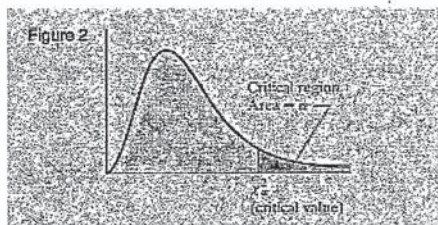
Step 3c:

Compute the test statistic $\chi_0^2 = \sum_{i=1}^k \frac{(O_i - E_i)^2}{E_i}$ where O_i is the *observed* count for category i .

Step 4 (Classical Approach):

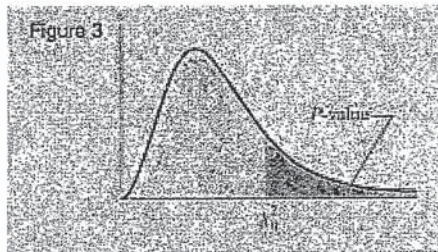
Determine the critical value using Table VIII.

All goodness-of-fit tests are right-tailed, so the critical value is χ_α^2 with $k - 1$ degrees of freedom (Figure 2). Compare the test statistic to the critical value. If $\chi_0^2 > \chi_\alpha^2$, reject the null hypothesis.



Step 4 (P-value Approach):

Use Table VIII or technology to approximate the P -value by determining the area under the chi-square distribution with $k - 1$ degrees of freedom to the **right** of the test statistic (Figure 3). If the P -value $< \alpha$, reject the null hypothesis.



Step 5: State the conclusion.

Remember, if we do *not* reject the null hypothesis, we are *not* saying that it *must* be true. We simply state that we do *not* have evidence to conclude it is as stated in the alternative hypothesis.

Instructions for
technology are given
on the next page.

Instructions for TI-84 Plus Calculators:

1. Enter the observed counts in **L1** and the expected counts in **L2** in the **STAT** editor.
2. Press **STAT > TESTS > D: χ^2 GOF-Test**.
3. Of course, tell it that **L1** and **L2**, respectively, are where the **Observed** and **Expected** counts are located. Calculate degrees of freedom as one less than the number of categories and enter it for **df**.
4. Highlight **Calculate** and press **ENTER**.
5. The calculator displays the test statistic, the *P*-value, and the degrees of freedom (**df**). The bottom line of output is labeled **CNTRB**. That shows the individual terms of the sum that makes up the test statistic. You can scroll through them with the right and left arrows.

TI-83 and older models of the TI-84 *cannot* do this test. However, a workaround can be found on the internet.

Instructions for StatCrunch:

1. Enter the observed counts in the first column and the expected counts in the second column. Name the columns.
2. Select **Stat > Goodness-of-fit > Chi-Square Test**.
3. Tell it the column that contains the **Observed** counts. Tell it the column that contains the **Expected** counts. Then hit **Compute!**.
4. Besides posting various bits for checking (like sample size, degrees of freedom, and the data), it will display the test statistic (**Chi-Square**) and the *P*-value.

expl 1: Is wearing a helmet on your motorcycle necessary? The National Highway Traffic Safety Administration publishes reports on motorcycle safety. The distribution shows the proportion of fatalities by location of injury for motorcycle accidents.

| Location of Injury | Multiple locations | Head | Neck | Thorax | Abdomen/Lumbar/Spine |
|--------------------|--------------------|------|------|--------|----------------------|
| Proportion | 0.57 | 0.31 | 0.03 | 0.06 | 0.03 |

The following data show the location of injury and number of fatalities for 2,068 riders who were not wearing a helmet.

| Location of Injury | Multiple locations | Head | Neck | Thorax | Abdomen/Lumbar/Spine |
|----------------------|--------------------|------|------|--------|----------------------|
| Count (out of 2,068) | 1036 | 864 | 38 | 83 | 47 |

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expl 1 (continued):

Use a 0.05 level of significance to test if the distribution of fatal injuries for riders not wearing a helmet follows the distribution for all riders. Follow these steps.

a.) Use the proportions given in the first table to calculate the expected counts for each category for the 2,068 riders who died while not wearing a helmet. Record them below. Do not round.

| Location of Injury | Multiple locations | Head | Neck | Thorax | Abdomen/Lumbar/Spine |
|-------------------------------|--------------------|--------|-------|--------|----------------------|
| Observed Count | 1036 | 864 | 38 | 83 | 47 |
| Assumed Proportion | 0.57 | 0.31 | 0.03 | 0.06 | 0.03 |
| Expected Count (out of 2,068) | 1178.76 | 641.08 | 62.04 | 124.08 | 62.04 |

b.) Write down the null and alternative hypotheses.

H_0 : The distribution for helmet-less riders is same as for all accidents

H_1 : The dist. is not the same for helmeted vs helmetless riders.

c.) Perform your Chi-Square test at the 0.01 level. Fill out the information below and state your conclusion.

Test statistic: $\chi^2_0 \approx 121.37$

P-value: $p \approx 2.7 \times 10^{-25} \approx 0$

Conclusion: $P\text{-value} < \alpha = 0.01$ — so reject null hyp. We have sufficient evidence to say the distribution for helmetless riders is not the same as for all riders.

d.) Our conclusion is not surprising, except possibly to those who do not believe in wearing helmets. Look through the CNTRB output (if you used the calculator) to find the category whose term contributed most to the test statistic. Is it surprising? (You can also compare observed and expected counts directly.)

multiple locations 17.3
head 77.5
neck 9.3
thorax 13.6
abdomen/spine 3.6

It's not surprising people w/out helmets hurt their head lots.

H_0 : All sides would have prob. of $\frac{1}{6}$

H_1 : H_0 is not true.

expl 2: I have made a six-sided die out of cardboard and must test it to see if it is fair. How would we do this?

Roll it, say, 100 times. Come up with sample data. For instance, it may look like

| (die roll) output | 1 | 2 | 3 | 4 | 5 | 6 |
|----------------------|----|----|----|----|----|----|
| trial out of 100 | 15 | 15 | 20 | 17 | 23 | 10 |

Compare these numbers to the expected counts. For each die number on a fair die, you'd expect $\frac{1}{6} \times 100 \approx 16.667$ rolls.

If we put those #'s in calculator for χ^2 , we get $\chi^2 \approx 6.08$ and a p-value of 0.29. So we'd not reject the null hypothesis at the $\alpha = 0.01, 0.05, \text{ or } 0.10$ levels.

Conclusion: We do not have sufficient evidence to say the die is not fair.

Table VIII

| Degrees of Freedom | Chi-Square (χ^2) Distribution Area to the Right of Critical Value | | | | | | | | | |
|--------------------|---|--------|--------|--------|--------|---------|---------|---------|---------|---------|
| | 0.995 | 0.99 | 0.975 | 0.95 | 0.90 | 0.10 | 0.05 | 0.025 | 0.01 | 0.005 |
| 1 | — | — | 0.001 | 0.004 | 0.016 | 2.706 | 3.841 | 5.024 | 6.635 | 7.879 |
| 2 | 0.010 | 0.020 | 0.051 | 0.103 | 0.211 | 4.605 | 5.991 | 7.378 | 9.210 | 10.597 |
| 3 | 0.072 | 0.115 | 0.216 | 0.352 | 0.584 | 6.251 | 7.815 | 9.348 | 11.345 | 12.838 |
| 4 | 0.207 | 0.297 | 0.484 | 0.711 | 1.064 | 7.779 | 9.488 | 11.143 | 13.277 | 14.860 |
| 5 | 0.412 | 0.554 | 0.831 | 1.145 | 1.610 | 9.236 | 11.070 | 12.833 | 15.086 | 16.750 |
| 6 | 0.676 | 0.872 | 1.237 | 1.635 | 2.204 | 10.645 | 12.592 | 14.449 | 16.812 | 18.548 |
| 7 | 0.989 | 1.239 | 1.690 | 2.167 | 2.833 | 12.017 | 14.067 | 16.013 | 18.475 | 20.278 |
| 8 | 1.344 | 1.646 | 2.180 | 2.733 | 3.490 | 13.362 | 15.507 | 17.535 | 20.090 | 21.955 |
| 9 | 1.735 | 2.088 | 2.700 | 3.325 | 4.168 | 14.684 | 16.919 | 19.023 | 21.666 | 23.589 |
| 10 | 2.156 | 2.558 | 3.247 | 3.940 | 4.865 | 15.987 | 18.307 | 20.483 | 23.209 | 25.188 |
| 11 | 2.603 | 3.053 | 3.816 | 4.575 | 5.578 | 17.275 | 19.675 | 21.920 | 24.725 | 26.757 |
| 12 | 3.074 | 3.571 | 4.404 | 5.226 | 6.304 | 18.549 | 21.026 | 23.337 | 26.217 | 28.306 |
| 13 | 3.565 | 4.107 | 5.009 | 5.892 | 7.042 | 19.812 | 22.362 | 24.736 | 27.688 | 29.819 |
| 14 | 4.075 | 4.660 | 5.629 | 6.571 | 7.790 | 21.064 | 23.685 | 26.119 | 29.141 | 31.319 |
| 15 | 4.601 | 5.229 | 6.262 | 7.261 | 8.547 | 22.307 | 24.996 | 27.488 | 30.578 | 32.801 |
| 16 | 5.142 | 5.812 | 6.908 | 7.962 | 9.312 | 23.542 | 26.296 | 28.845 | 32.000 | 34.267 |
| 17 | 5.697 | 6.408 | 7.564 | 8.672 | 10.085 | 24.769 | 27.587 | 30.191 | 33.409 | 35.718 |
| 18 | 6.265 | 7.015 | 8.231 | 9.390 | 10.865 | 25.989 | 28.869 | 31.526 | 34.805 | 37.156 |
| 19 | 6.844 | 7.633 | 8.907 | 10.117 | 11.651 | 27.204 | 30.144 | 32.852 | 36.191 | 38.582 |
| 20 | 7.434 | 8.260 | 9.591 | 10.851 | 12.443 | 28.412 | 31.410 | 34.170 | 37.566 | 39.997 |
| 21 | 8.034 | 8.897 | 10.283 | 11.591 | 13.240 | 29.615 | 32.671 | 35.479 | 38.932 | 41.401 |
| 22 | 8.643 | 9.542 | 10.982 | 12.338 | 14.041 | 30.813 | 33.924 | 36.781 | 40.289 | 42.796 |
| 23 | 9.260 | 10.196 | 11.689 | 13.091 | 14.843 | 32.007 | 35.172 | 38.076 | 41.638 | 44.181 |
| 24 | 9.886 | 10.856 | 12.401 | 13.848 | 15.659 | 33.196 | 36.415 | 39.364 | 42.980 | 45.559 |
| 25 | 10.520 | 11.524 | 13.120 | 14.611 | 16.473 | 34.382 | 37.652 | 40.646 | 44.314 | 46.928 |
| 26 | 11.160 | 12.198 | 13.844 | 15.379 | 17.292 | 35.563 | 38.885 | 41.923 | 45.642 | 48.290 |
| 27 | 11.808 | 12.879 | 14.573 | 16.151 | 18.114 | 36.741 | 40.113 | 43.195 | 46.963 | 49.645 |
| 28 | 12.461 | 13.565 | 15.308 | 16.928 | 18.939 | 37.916 | 41.337 | 44.461 | 48.278 | 50.993 |
| 29 | 13.121 | 14.256 | 16.047 | 17.708 | 19.768 | 39.087 | 42.557 | 45.722 | 49.588 | 52.336 |
| 30 | 13.787 | 14.953 | 16.791 | 18.493 | 20.599 | 40.256 | 43.773 | 46.979 | 50.892 | 53.672 |
| 40 | 20.707 | 22.164 | 24.433 | 26.509 | 29.051 | 51.805 | 55.758 | 59.342 | 63.691 | 66.766 |
| 50 | 27.991 | 29.707 | 32.357 | 34.764 | 37.689 | 63.167 | 67.565 | 71.420 | 76.154 | 79.490 |
| 60 | 35.534 | 37.485 | 40.482 | 43.188 | 46.459 | 74.397 | 79.082 | 83.298 | 88.579 | 91.952 |
| 70 | 43.275 | 45.442 | 48.758 | 51.739 | 55.329 | 85.527 | 90.531 | 95.023 | 100.425 | 104.215 |
| 80 | 51.172 | 53.540 | 57.153 | 60.391 | 64.278 | 96.578 | 101.879 | 106.629 | 112.329 | 116.321 |
| 90 | 59.196 | 61.754 | 65.647 | 69.126 | 73.291 | 107.565 | 113.145 | 118.136 | 124.116 | 128.299 |
| 100 | 67.328 | 70.065 | 74.222 | 77.929 | 82.358 | 118.498 | 124.342 | 129.561 | 135.807 | 140.169 |

